

# **Decomposing Productivity Growth in the U.S. Computer Industry\***

**Hyunbae Chun\*\***

*Queens College  
City University of New York*

**M. Ishaq Nadiri\*\*\***

*New York University  
National Bureau of Economic Research*

**December 5, 2005**

## **Abstract**

In this paper, we examine the sources of the productivity growth in the U.S. computer industry from 1978 to 1999. We estimate a joint production model of output quantity and quality that distinguishes two types of technological changes: process and product innovations. Based on the estimation results, we decompose total factor productivity (TFP) growth rate into the contributions of process and product innovations and scale economies. The results show that product innovation associated with better quality accounts for about 30 percent of the TFP growth in the computer industry. Furthermore, the TFP acceleration in the computer industry in the late 1990s is mainly derived from a rapid increase in product innovation.

*Keywords:* Computer, Productivity, Process and Product Innovation, Hedonic Price  
*JEL Classification Numbers:* D24, L63, O33

---

\* We are especially grateful to Daron Acemoglu (the editor) and two anonymous referees for very helpful comments and insights. We also thank Diego Comin, Boyan Jovanovic, and Edward Wolff for their suggestions and Steve Rosenthal for providing the data used in this paper. The financial support of the C.V. Starr Center for Applied Economics at New York University and the PSC-CUNY at the City University of New York is also gratefully acknowledged.

\*\* Department of Economics, Queens College, City University of New York, 65-30 Kissena Blvd., Flushing, NY 11367 Tel: (718) 997-5450 Fax: (718) 997-5466 E-mail: hchun@qc1.qc.edu

\*\*\* Corresponding author. Department of Economics, New York University, 269 Mercer Street, 7th Floor, New York, NY 10003 Tel: (212) 998-8968 Fax: (212) 995-4186 E-mail: min1@nyu.edu

# 1. Introduction

During the last few decades, there has been a remarkable productivity growth in the production of information technology (IT) products such as computers, communications equipment, and semiconductors. A typical measure of productivity is total factor productivity (TFP), defined as the amount of output produced from a given amount of input. Hence, the traditional TFP approach mainly focuses on how much productivity growth is caused by the improvement in the technological efficiency of production process (process innovation).

In contrast to process innovation, productivity growth can take place in the improvement of output quality (product innovation). In particular, improvement in output quality is one of the most prevailing characteristics in IT production such as microprocessor speed, the capacity of storage devices and memory, etc. This suggests that technological innovation associated with better quality can be an important source of the TFP growth in the IT-producing industry. As Hulten (2001) pointed out, however, the TFP approach is silent about product innovation.<sup>1</sup> Therefore, the identification of both process and product innovations is crucial to the exploration of the sources of productivity growth in the IT-producing industry.

In this paper, we examine the sources of the productivity growth in the U.S. computer industry from 1978 to 1999. The novelty of this paper is that we separate two different technical changes in TFP growth: product innovation associated with better quality and process innovation associated with more quantity. Using both the hedonic

---

<sup>1</sup> Although some recent studies by Jorgenson and Stiroh (2000), Oliner and Sichel (2000), and Whelan (2002) have attempted to measure the TFP growth in the IT-producing industry, there have been few studies that distinguish the contributions of process and product innovations in the productivity growth in this industry.

(quality-adjusted) and list (quality-unadjusted) prices, we construct the variables of output quantity and quality. Then, we formulate the joint production model of output quantity and quality, and estimate the joint optimization conditions of quantity and quality together with a general cost structure that accounts for scale economies and markups. Based on estimation results, we find that technical change associated with process and product innovations contributes about 70 percent of TFP growth in the computer industry, while the effect of scale economies explains about 30 percent of the TFP growth. In particular, product-oriented technical change explains about 30 percent of the TFP growth in the computer industry. Furthermore, we find that the TFP contribution from product innovation rapidly rose in the late 1990s, while the contribution from process innovation and economies of scale changed little.

## 2. Measuring Output Quality and Quantity

Product innovation takes place in the improvement of quality or in the introduction of new products that have better quality. Thus, measuring changes in output quality, given physical units of output, is crucial to distinguish product innovation from process innovation. The hedonic price method provides us a viable solution to measure the improvement of output quality. The hedonic price corrects the list price (unit price) of output for changes in product attributes that characterize changes in output quality. The quality of output ( $Y_s$ ) is given by

$$Y_s = \left( \frac{P}{P_Y} \right) \quad (1)$$

where  $P$  is the list price of output and  $P_Y$  is the hedonic price of output. It is important to note that drops in the hedonic price can be greater than true quality improvement if the list price falls due to changes in the demand factors such as declines in markup.<sup>2</sup>

The quantity of output (physical units of output) can be obtained from dividing the nominal value of production ( $V_Y$ ) by the list price of output,

$$Y_Q = \left( \frac{V_Y}{P} \right) \quad (2)$$

where  $Y_Q$  is the quantity of output. The quality-adjusted output can be obtained from multiplying output quantity by output quality.

In this paper, we constructed output quantity and quality in four computer industries: electronic computers (3571), computer storage devices (3572), computer terminals (3575), and computer peripheral equipment (3577) (1987 SIC codes in parentheses). Using the ratio of the list price to the hedonic price, we constructed the output quality. We obtained the hedonic price for computer products from the Bureau of Economic Analysis (BEA).<sup>3</sup>

The list price was obtained from the *Current Industrial Reports (CIR)* published by the Census Bureau. The CIR includes the nominal values and physical units of shipments for detailed products in each industry. For example, the electronic computers industry in 1997 has 11 types of computers such as host and single-user computers such as large- and medium-scale systems, PC servers, personal computers, workstations,

---

<sup>2</sup> The hedonic prices capture the relationship between equilibrium price and product quality, not just the valuation of the quality. The importance of the market condition in the hedonic price method is also emphasized in the studies by Rosen (1974) and Pakes (2003). Imperfect competition leads to markups being positively correlated with quality so that the hedonic price can overestimate improvement in quality (Hobijn, 2001).

notebooks, and so on. We first construct the list price of each product dividing the nominal value of shipments by the physical units of shipments. To avoid the bias in the list price that can result from changes in the composition of products, we use the Törnqvist index to construct the industry-level list price.<sup>4</sup> The growth rate of list price at the industry-level is calculated as the weighted average of the growth rates of all products in an industry where the weight of each product is defined as an average nominal value of shipments over two adjacent years.

[Table 1]

Table 1 presents the annual growth rates of hedonic and list prices and output quality and quality for the period 1978-1999. As shown in equation (1), the growth rate of quality is defined as the growth rate of the list price minus the growth rate of the hedonic price. The annual growth rate of the quality improvement is about 12 percent on average for all industries from 1978 to 1999. This suggests that the improvement in true quality is smaller than changes in the hedonic price. Quality improvements are significantly different across industries. The quality improvement in electronic computers has been faster than those in the other three computer products. The average annual growth rate of quantity is about 15 percent for all industries. The growth rate of the quality-adjusted output, that is the sum of the growth rates of quantity and quality, is

---

<sup>3</sup> See BEA (1998) for the details on sources and methods of the hedonic price index.

<sup>4</sup> We thank an anonymous referee for suggesting us to correct for these biased results from changes in the mix of products. For example, if the share of cheaper products increases, the industry-level list price can decline without changes in each product's list price. In fact, most industries (except for computer peripheral equipment) show this pattern. Therefore, without correcting for the compositional changes, declines in the list price can be exaggerated and subsequently, the quality improvement can be underestimated.

about 27 percent. This implies that quality improvement contributes almost a half of the growth of the quality-adjusted output.

### 3. Empirical Specification

A firm chooses both output quantity and quality to maximize its profits,

$$\pi(Y_Q, Y_S) = Y_Q P(Y_Q, Y_S) - C(Y_Q, Y_S) \quad (3)$$

where  $P(Y_Q, Y_S)$  is the inverse demand function and  $C(Y_Q, Y_S)$  is the cost function. The quality improvement shifts the quantity demand curve to the right and raises the curve of cost. Thus, the optimal choice of output quality depends both on the demand elasticity for quality and on the marginal cost of quality production.<sup>5</sup>

Main purpose of our model is to quantify the contribution of both product and process innovations to productivity growth.<sup>6</sup> Thus, we are interested in measuring the marginal costs of both the quality and quantity that are determined by all factors of production. To accomplish this, we use a flexible cost function that allows a general cost structure of quality and quantity production. In contrast, the quality improvement in the endogenous growth models of Grossman and Helpman (1991) and Aghion and Howitt (1992) is determined only by specific inputs such as research and development (R&D) and skilled workers.<sup>7</sup>

---

<sup>5</sup> If the cost structure for quality production has a property of constant returns to scale, the choice of quality production is dependent upon the demand condition, but is independent of the cost structure.

<sup>6</sup> Athey and Schmutzler (1995) emphasize complementarities between product (demand-enhancing) and process (cost-saving) innovations, which suggests that the two innovations cannot be separately determined in the model.

<sup>7</sup> As shown in Scherer (1984) suggesting that more than 90 percent of R&D in the computer industry is devoted to product innovation, investment in R&D can be an important source of product innovation. Since

We specify a translog cost function to describe the technology of the firm as:

$$\begin{aligned}
\ln(c_t^v) = & \beta_0 + \beta_L \ln w_t + \beta_K \ln K_{t-1} + \sum_{j=Q,S} \beta_j \ln Y_{j,t} + \beta_T T_t \\
& + \frac{1}{2} \left[ \beta_{LL} (\ln w_t)^2 + \beta_{KK} (\ln K_{t-1})^2 + \sum_{j=Q,S} \beta_{jj} (\ln Y_{j,t})^2 + \beta_{TT} T_t^2 \right] \\
& + \beta_{LK} \ln w_t \ln K_{t-1} + \sum_{j=Q,S} \beta_{Lj} \ln w_t \ln Y_{j,t} + \beta_{LT} \ln w_t T_t \\
& + \sum_{j=Q,S} \beta_{Kj} \ln K_{t-1} \ln Y_{j,t} + \beta_{KT} \ln K_{t-1} T_t \\
& + \beta_{QS} \ln Y_{Q,t} \ln Y_{S,t} + \sum_{j=Q,S} \beta_{jT} \ln Y_{j,t} T_t
\end{aligned} \tag{4}$$

where the variable cost is given by  $C^v = W_L L + W_M M$  and the variable cost and the price of labor input are normalized by the price of materials, i.e.,  $c^v = (C^v/W_M)$  and  $w = (W_L/W_M)$ , respectively. The normalization imposes the homogeneity restriction on the cost function.  $K_{t-1}$  is a lagged variable of quasi-fixed capital stock and  $T$  is an index of process-oriented technical change that represents shift in the variable cost function.

Rearranging two first-order conditions for profit maximization with respect to output quantity and quality with the translog cost function, we can derive the variable cost share equations of output quantity and quality as:

$$S_{Q,t} = (1 + \mu_Q) \frac{\partial \ln C_t^v}{\partial \ln Y_{Q,t}} = \left( \frac{1}{1 + \alpha_Q} \right) (\beta_Q + \beta_{LQ} \ln w_t + \beta_{KQ} \ln K_{t-1} + \sum_{j=Q,S} \beta_{Qj} \ln Y_{j,t} + \beta_{QT} T_t) \tag{5}$$

$$S_{S,t} = (1 + \mu_S) \frac{\partial \ln C_t^v}{\partial \ln Y_{S,t}} = \left( \frac{1}{\alpha_S} \right) (\beta_S + \beta_{LS} \ln w_t + \beta_{KS} \ln K_{t-1} + \sum_{j=Q,S} \beta_{jS} \ln Y_{j,t} + \beta_{ST} T_t) \tag{6}$$

---

we consider both the product and process innovations, adopting R&D in our model also requires separate R&D expenditures on the product and process innovations. This is beyond the scope of our study and will be addressed in future research.

where  $S_Q$  is the variable cost share of output quantity,  $\mu_Q$  is the markup for output quantity, and  $\alpha_Q$  is the inverse demand elasticity with respect to output quantity.  $S_S$ ,  $\mu_S$ , and  $\alpha_S$  are similarly defined for the quality of output. Since  $\alpha_S$  represents the willingness to pay for one more unit of quality, the supply of the better quality can increase the revenue by way of raising the price. Thus, we can expect the positive sign for  $\alpha_S$  in contrast to the negative sign for  $\alpha_Q$ .

Aghion and Howitt (1992) focus on the growth implication of the creative destruction due to the quality innovation that a new innovation creates monopoly rents, but destroys rents from the previous innovation. In this regard, they assume that the product quality chosen by innovating firms is at the frontier. But, the optimal quality in our model is not necessarily at the frontier level, which reflects that some firms within the industry produce computers at the frontier level, but others not. For example, Bresnahan, Stern, and Trajtenberg (1997) find that both frontier- and non-frontier-level personal computers coexist at the same time.

Applying Shephard's lemma to the variable cost function, we derive the cost share equations of labor ( $S_L = W_L L / C^v$ ). The variable cost share of materials can be derived as  $S_M = I - S_L$ . Using the envelope theorem, we can derive the long-run equilibrium condition for a quasi-fixed factor of capital. In the long-run equilibrium, the optimal quantity of the capital stock is determined by the condition that the rental rate of capital is equal to the magnitude of the reduction in variable cost due to an increase in an additional unit of the capital stock. This condition yields the variable cost share equation of the capital stock ( $S_K = W_K K / C^v$ ) where  $W_K$  is the user cost of capital.

The growth rate of TFP is traditionally defined as the difference between the growth rate of output and the growth rate of all inputs. Since the growth rate of the quality-adjusted output is equal to the sum of growth rates of output quantity and quality, the rate of TFP growth in a two-output case can be given by

$$TFP_t = \sum_{j=Q,S} \dot{Y}_{j,t} - \sum_{i=L,M,K} \frac{1}{2} (\tilde{S}_{i,t-1} + \tilde{S}_{i,t}) \dot{X}_{i,t} \quad (7)$$

where a dot over the variable denotes the rate of growth,  $\tilde{S}_i$  is the total cost share of input  $i$ , and  $\dot{X}_i$  is the growth rate of input  $i$ .

Following the methodology of the TFP decomposition in a multiple-output case proposed by Denny, Fuss, and Waverman (1981) and Nadiri and Nandi (1999), we can decompose the growth rate of TFP into three factors as:

$$TFP = (1 - \frac{1}{\rho_Q}) \dot{Y}_Q + (1 - \frac{1}{\rho_S}) \dot{Y}_S - \eta_T \quad (8)$$

$$\rho_j = \frac{1 - \eta_K^v}{\eta_j^v}, \quad \eta_K^v = \frac{\partial \ln C^v}{\partial \ln K}, \quad \eta_j^v = \frac{\partial \ln C^v}{\partial \ln Y_j}, \quad \eta_T = \frac{\eta_T^v}{1 - \eta_K^v}, \quad \text{and} \quad \eta_T^v = \frac{\partial \ln C^v}{\partial T} \quad \text{for } j = Q, S$$

where  $\eta_K^v$  is the variable cost elasticity with respect to the capital stock,  $\eta_Q^v$  and  $\eta_S^v$  are the variable cost elasticities with respect to output quantity and quality, respectively, and  $\eta_T^v$  is the variable cost elasticity with respect to the time variable. Scale economies are measured as the inverse of cost elasticity with respect to output quantity. After correcting for the effect of a quasi-fixed factor of the capital stock on the variable cost,  $\rho_Q$  measures scale economies in the long-run. There are economies of scale if  $\rho_Q$  is greater than one.

In a similar vein,  $\rho_s$  measures the degree of the cost efficiency for product innovation which can be a source of productivity growth if  $\rho_s$  is greater than one. Process-oriented technical change can be measured with  $-\eta_T$  that represents a shift in the cost function. On the right-hand side of equation (8), the first term is scale effect, the second term is the effect of product innovation, and the last term is the effect of process innovation.<sup>8</sup>

## 4. Results

We estimate a system of equations consisting of the variable cost function, the variable cost share equations of labor and capital, and two optimality conditions for output quantity and quality using the non-linear three-stage least squares (3SLS).<sup>9</sup> We use a set of instrumental variables: lagged variables and macroeconomic variables such as oil price, defense spending, the number of population, GDP per capita, the share of nonmanufacturing sector, and corporate income tax rate.<sup>10</sup> We also estimated the model with various sets of instruments. In particular, lagged variables are not good instruments if the error terms are autocorrelated. Thus we corrected for first-order autocorrelation in the error terms and estimated the model without lagged variables, but the elasticity estimates are qualitatively not different.

---

<sup>8</sup> Nadiri and Nandi (1999) decomposed scale effect into several exogenous components such as changes in exogenous demands, factor prices, etc. Unless either quantity or quality is assumed to be exogenous, we cannot decompose the scale and quality effects into other exogenous factors because of the property of the joint determination of quantity and quality.

<sup>9</sup> In addition to output quantity and quality, we use the *NBER-CES Manufacturing Industry Database* (Bartelsman and Gray, 1996) to construct the quantity and price of labor, capital, and materials. Since the NBER-CES database covers only up to 1996, so we expanded the database to 1999 with the *Annual Survey of Manufactures* and the *Census of Manufactures*.

<sup>10</sup> Since computers are more intensively used in the nonmanufacturing sector than the manufacturing, we use the nonmanufacturing share as an instrument for the demand for computers in the U.S. business sector. To avoid endogeneity in GDP per capita, we also excluded the value-added of the computer industry from GDP.

The optimal choice of output quantity and quality depends on two different demand conditions as well as cost structures that are jointly estimated in this study. We introduce the industry dummy variables for intercepts in all equations, and also allow the industry dummy variables for time coefficient to capture different process innovations among the four industries. The majority of the parameter estimates are statistically significant.<sup>11</sup>

### **Cost Elasticities and TFP Growth Decomposition**

Table 2 presents the short-run and long-run cost elasticities with respect to output quantity and quality, input prices, a quasi-fixed factor of the capital stock, and time variable. The short-run implies that production is conducted when the level of the capital stock is fixed, while in the long-run, cost can be minimized with the adjustment of capital.

Since elasticity estimates are non-linear functions of parameters, the standard errors are obtained by using the bootstrap method with replacement 1000 times.<sup>12</sup> We also calculated the standard errors using the delta method that uses a first-order variance approximation under a normality assumption. Both the bootstrap and delta methods' standard errors are very close to each other.

[Table 2]

In Table 2, an interesting result is that the cost elasticity of output quality is smaller than that of output quantity. This suggests that the production of quality entails

---

<sup>11</sup> The parameter estimates are available from the authors upon request.

less cost than the production of quantity. Producing better quality goods is not subject to sharply increasing cost if there are learning-by-doing and spillovers of new technologies. For example, Irwin and Klenow (1994) find that learning-by-doing is an important source of cost reduction in semiconductor production. The variable cost elasticity of the time variable suggests that the variable cost has declined about 6.7 percent annually, given production of the same amount of output quantity and quality.<sup>13</sup>

The bottom panel of Table 2 presents the decomposed contribution to TFP growth that includes the effects of process and product innovations and scale effect. An average TFP growth rate for the four industries is about 15 percent, but the TFP growth rates vary considerably across the industries. For example, the TFP growth rate in the electronic computers industry is approximately 18 percent while the in the terminals industry is only one-third of the rate in the electronic computers industry.

The contribution of product innovation is about 4.2 percent, which accounts for about 30 percent of TFP growth in the computer industries. Process innovation explains about 40 percent of TFP growth in these industries, while scale effect does about 30 percent. The sum of process and product innovations in total explains about 70 percent of the TFP growth, which implies that the rapid TFP growth in the U.S. computer industries is mainly attributable to fast improvement in technologies.

Our estimate of TFP contribution from product innovation can be downward biased if the BEA hedonic price underestimates improvement in quality. Recently, Pakes

---

<sup>12</sup> See Efron and Tibshirani (1993) for the details on the bootstrap standard error and Eakin, McMillen, and Buono (1990) for its applications in the standard errors of elasticity estimates in a cost function.

<sup>13</sup> In addition to cost elasticities, the estimates of two demand side parameters of  $\alpha_Q$  and  $\alpha_S$  are -0.387 and 0.519, respectively. This implies that the demand elasticity of quantity ( $1/\alpha_Q$ ) is higher than that of quality ( $1/\alpha_S$ ). The finding also suggests that the computer product market is more competitive in quantity than in quality.

(2003) pointed out possible biases in the official hedonic indexes resulted from inappropriate sample selection correction for the high attrition rate of computers as well as estimation variances of hedonic indexes. In a similar vein, Berndt, Dulberger, and Rappaport (2001) also show that the hedonic price estimation is very sensitive to the choice of model specification. In spite of a possible underestimated quality in the BEA hedonic price, our estimate provides at least a lower bound of product innovation.<sup>14</sup>

### **Implications for Productivity Growth in the Aggregate Economy**

There was a remarkable resurgence in the productivity growth from the early to late 1990s. Determining the source of the productivity revival in the aggregate economy, some recent studies (Jorgenson and Stiroh, 2000; Oliner and Sichel, 2000) focus on the role of the IT-producing industry. Table 3 shows that the contribution from the computer industry to the aggregate TFP growth is on average 0.22 that accounts for almost one-third of the aggregate TFP growth.<sup>15</sup> Moreover, the TFP contribution of the computer industry has risen about two times from the early to late 1990s. The TFP acceleration in the computer industry from the early to late 1990s explains about 40 percent of the acceleration in the aggregate TFP growth.

[Table 3]

Our approach contrasts the conventional TFP accounting in two ways. We explicitly consider the effect of product demand and economies of scale, which enables

---

<sup>14</sup> We thank two anonymous referees for suggesting to us the possible biases in the hedonic prices.

<sup>15</sup> We also compare our TFP growth rate in the computer industry with that of Oliner and Sichel (2000) measured by the dual approach. The TFP growth rates are different between this study and their study during the second half of the 1990s. One of the main reasons for the difference is that they assume the same cost structure between the IT-producing industry and other industries.

us to identify non-technological factors in the TFP growth. In contrast, Oliner and Sichel (2000) and Whelan (2002) assume constant returns to scale and perfect competition, and thus may overestimate the degree of technological change in the industry.<sup>16</sup> In fact, we find that the scale effect explains almost 30 percent of the contribution suggesting that this contribution is associated with non-technological factors.

Second, we identify the contribution of product innovation in the computer industry to TFP growth in the aggregate economy, but the conventional TFP accounting models do not. As Grossman and Helpman (1991) pointed out, quality improvement has played a central role in economic growth. However, studies measuring the contribution of quality improvement in economic growth are rare. Focusing on the computer industry experiencing rapid improvements in quality, we try to quantify the contribution of product innovation in this industry to the aggregate productivity growth.

Although the contribution of process innovation to the TFP growth in the aggregate economy is greater than that of product innovation, an increase in the product-oriented technical change in the late 1990s explains about 60 percent of the acceleration in the computer industry's TFP contribution. The contribution from the non-technological factor of scale economies also changed little from the early to late 1990s. Therefore, the findings suggest that the productivity acceleration in the aggregate economy associated with the computer industry during the late 1990s is largely attributable to acceleration in the product innovation in this industry.

Furthermore, Table 3 shows that the contribution of product innovation rose in the sample period while the contribution from process innovation and economies of scale

---

<sup>16</sup> See Nadiri and Prucha (2001) more detailed explanations on possible biases in measuring TFP due to economies of scales and imperfect competition.

changed little. Filson (2001) also finds an increasing trend in product innovation in the computer industry, but a decreasing trend in the automobile industry. The time-series pattern of product and process innovations in the computer industry challenges the prediction of the industry life-cycle theory that new industries experience product innovation early and process innovation later (Gort and Klepper, 1982; Cohen and Klepper, 1996).

## **5. Conclusion**

In this paper, we provide an empirical framework for exploring the different sources of the productivity growth in the U.S. computer industry. The empirical results of this study show that technological changes associated with both process and product innovations have been a major source of the TFP growth in the computer industry. Furthermore, we quantify the contribution of quality improvement in computers to the productivity growth in the U.S. economy. We find that a rapid increase in the quality of computers has a significant contribution to the TFP acceleration of the U.S. economy in the late 1990s.

## References

Aghion, Philippe and Peter Howitt, "A Model of Growth through Creative Destruction," *Econometrica*, 60(2), March 1992, 323-351.

Athey, Susan and Armin Schmutzler, "Product and Process Flexibility in an Innovative Environment," *Rand Journal of Economics*, 26(4), Winter 1995, 557-574.

Bartelsman, Eric J. and Wayne Gray, "The NBER Productivity Database," NBER Technical Working Paper No. 205, October 1996.

Berndt, Ernst R., Ellen R. Dulberger, and Neal J. Rappaport, "Price and Quality of Desktop and Mobile Computers: A Quarter Century Historical Overview," *American Economic Review*, 91(2), May 2001, 268-273.

Bresnahan, Timothy F., Scott Stern, and Manuel Trajtenberg, "Market Segmentation and the Sources of Rents from Innovation: Personal Computers in the Late 1980s," *Rand Journal of Economics*, 28, 1997, S17-S44.

Bureau of Economic Analysis, "Computer Prices in the National Accounts: An Update from the Comprehensive Revision," National Income and Wealth Division Working Paper, Bureau of Economic Analysis, August 1998.

Bureau of Labor Statistics, *Multifactor Productivity Trends, 1999*, Bureau of Labor Statistics, May 2001.

Cohen Wesley M. and Steven Klepper, "Firm Size and the Nature of Innovation within Industries: The Case of Process and Product R&D," *Review of Economics and Statistics*, 78(2), May 1996, 232-243.

Denny, Michael, Melvyn Fuss, and Leonard Waverman "The Measurement and Interpretation of Total Factor Productivity in Regulated Industries with an Application to Canadian Telecommunications," in Thomas G. Cowing and Rodney E. Stevenson, (eds.), *Productivity Measurement in Regulated Industries*, New York: Academic Press, 1981, 179-218.

Eakin, Kelly B., Daniel P. McMillen, and Mark J. Buono, "Constructing Confidence Intervals Using the Bootstrap: An Application to a Multi-Product Cost Function," *Review of Economics and Statistics*, 72(2), May 1990, 339-344.

Efron, Bradley and Robert J. Tibshirani, *An Introduction to the Bootstrap*, New York: Chapman & Hall/CRC, 1993.

Filson, Darren, "The Nature and Effects of Technological Change over the Industry Life Cycle," *Review of Economic Dynamics*, 4(2), 2001, 460-494.

Gort, Michael and Steven Klepper, "Time Paths in the Diffusion of Product Innovations," *Economic Journal*, 92(367), September 1982, 630-653.

Grossman, Gene M. and Elhanan Helpman, "Quality Ladders in the Theory of Growth," *Review of Economic Studies*, 58(1), January 1991, 43-61

Hobijn, Bart, "Is Equipment Price Deflation a Statistical Artifact?" Federal Reserve Bank of New York Staff Reports No. 139, November 2001.

- Hulten, Charles R. "Total Factor Productivity: A Short Bibliography," in Charles R. Hulten, Edwin R. Dean, and Michael J. Harper (eds.), *New Developments in Productivity Analysis*, Chicago: University of Chicago Press, 2001, 1-47.
- Irwin, Douglas A. and Peter J. Klenow, "Learning-by-Doing Spillovers in the Semiconductor Industry," *Journal of Political Economy*, 102(6), December 1994, 1200-1227.
- Jorgenson, Dale W. and Kevin J. Stiroh, "Rising the Speed Limit: U.S. Economic Growth in the Information Age," *Brookings Papers on Economic Activity*, (1), 2000, 125-233.
- Nadiri, M. Ishaq and Banani Nandi, "Technical Change, Markup, Divestiture, and Productivity Growth in the U.S. Telecommunications Industry," *Review of Economics and Statistics*, 81(3), August 1999, 488-498.
- Nadiri, M. Ishaq and Ingmar R. Prucha, "Dynamic Factor Demand Models and Productivity Analysis," in Charles R. Hulten, Edwin R. Dean, and Michael J. Harper (eds.), *New Developments in Productivity Analysis*, Chicago: University of Chicago Press, 2001, 103-164.
- Oliner, Stephen D. and Daniel E. Sichel, "The Resurgence of Growth in the Late 1990s: Is Information Technology the Story?" *Journal of Economic Perspectives*, 14(4), Fall 2000, 3-22.
- Pakes, Ariel, "A Reconsideration of Hedonic Price Indexes with an Application to PC's," *American Economic Review*, 93(5), December 2003, 1578-1596.
- Rosen, Sherwin, "Hedonic Prices and Implicit Markets: Product Differentiation in Pure Competition," *Journal of Political Economy*, 82(1), January-February 1974, 34-55.
- Scherer, Frederic M., "Using Linked Patent and R&D Data to Measure Interindustry Technology Flows," in Zvi Griliches (ed.) *R&D, Patents, and Productivity*, Chicago: University of Chicago Press, 1984, 417-461.
- Whelan, Karl, "Computers, Obsolescence, and Productivity," *Review of Economics and Statistics*, 84(3), August 2002, 445-461.

**Table 1. Hedonic and List Prices and Output Quality and Quantity, 1978-1999<sup>a</sup>**

	Electronic computers (3571)	Computer storage devices (3572)	Computer terminals (3575)	Computer peripheral equipment (3577)	Average
Hedonic price	-20.469	-8.586	-8.370	-12.848	-16.873
List price	-5.214	-0.477	-1.304	-8.930	-5.228
Quality of output	15.255	8.108	7.066	3.918	11.645
Quantity of output	15.411	11.553	2.149	18.101	15.134

Note: Average growth rates of the four industries are weighted by the industry share of nominal output between two adjacent years. The growth rate of output quantity is equal to the growth rate of nominal output minus the growth rate of list price. The growth rate of output quality is equal to the growth rate of list price minus the growth rate of hedonic price.

<sup>a</sup>: percentage per year

**Table 2. Cost Elasticities and TFP Growth Decomposition**

	Electronic computers (3571)	Computer storage devices (3572)	Computer terminals (3575)	Computer peripheral equipment (3577)	Average
<i>Short-run Elasticities</i>					
Quantity of output	0.829 (0.040)	0.730 (0.036)	0.705 (0.034)	0.713 (0.035)	0.787 (0.038)
Quality of output	0.702 (0.041)	0.619 (0.037)	0.597 (0.035)	0.604 (0.036)	0.666 (0.039)
Time	-0.090 (0.008)	-0.002 (0.005)	-0.022 (0.003)	-0.049 (0.007)	-0.067 (0.006)
Price of labor input	0.232 (0.013)	0.305 (0.013)	0.311 (0.013)	0.286 (0.012)	0.255 (0.008)
Price of materials	0.768 (0.013)	0.695 (0.013)	0.689 (0.013)	0.714 (0.012)	0.745 (0.008)
Capital stock	-0.093 (0.013)	-0.120 (0.012)	-0.108 (0.013)	-0.073 (0.012)	-0.092 (0.008)
Scale	1.206 (0.063)	1.369 (0.072)	1.418 (0.074)	1.402 (0.075)	1.271 (0.066)
<i>Long-run Elasticities</i>					
Quantity of output	0.759 (0.038)	0.652 (0.033)	0.636 (0.031)	0.664 (0.033)	0.720 (0.035)
Quality of output	0.643 (0.038)	0.552 (0.033)	0.539 (0.033)	0.563 (0.033)	0.610 (0.036)
Time	-0.082 (0.007)	-0.002 (0.005)	-0.020 (0.003)	-0.046 (0.007)	-0.061 (0.006)
Price of labor input	0.212 (0.012)	0.272 (0.012)	0.281 (0.012)	0.266 (0.012)	0.234 (0.008)
Price of materials	0.703 (0.015)	0.621 (0.013)	0.621 (0.013)	0.665 (0.014)	0.682 (0.010)
Scale	1.318 (0.072)	1.534 (0.084)	1.572 (0.082)	1.505 (0.082)	1.389 (0.074)
<i>TFP Growth Decomposition<sup>a</sup></i>					
TFP growth rate	18.463	7.999	5.996	10.905	15.059
Scale effect	3.755 (0.592)	4.078 (0.378)	0.681 (0.066)	6.104 (0.605)	4.255 (0.551)
Effect of product innovation	5.295 (0.579)	3.555 (0.269)	3.300 (0.231)	1.650 (0.131)	4.230 (0.413)
Effect of process innovation	8.193 (0.716)	0.199 (0.471)	2.013 (0.276)	4.600 (0.680)	6.147 (0.578)
Residuals	1.221 (0.272)	0.166 (0.244)	0.002 (0.218)	-1.448 (0.266)	0.427 (0.182)
<i>Industry Output Share</i>	0.616	0.130	0.034	0.220	

Note: Numbers in parentheses are the standard errors that are evaluated at their mean values and are also obtained by using the bootstrap method with replacement 1000 times. Average elasticities, scales, and decomposed effects of the four industries are weighted by the industry share of nominal output.

<sup>a</sup>: percent per year.

**Table 3. Contribution of the Computer Industry to TFP Growth in the Non-farm Business Sector**

	1978-1990	1991-1995	1996-1999	1978-1999
<i>TFP growth rate in the non-farm business sector:</i> <sup>a</sup>				
BLS (2001)	0.303	0.603	1.098	0.516
<i>TFP growth rate in the computer industry:</i> <sup>a</sup>				
Oliner and Sichel (2000)	11.2 <sup>b</sup>	11.3	16.6	12.2 <sup>c</sup>
This study	13.98	12.37	22.05	15.06
<i>Domar weight of the computer industry:</i>				
Oliner and Sichel (2000)	0.011 <sup>b</sup>	0.014	0.016	0.012 <sup>c</sup>
This study	0.014	0.014	0.017	0.015
<i>Contribution from the computer industry to TFP growth in the non-farm business sector:</i> <sup>a</sup>				
Oliner and Sichel (2000)	0.12 <sup>b</sup>	0.16	0.26	0.15 <sup>c</sup>
This study				
Total contribution	0.193	0.170	0.370	0.220
Scale effect	0.070	0.052	0.024	0.058
	(0.009)	(0.007)	(0.005)	(0.008)
Effect of product innovation	0.037	0.046	0.171	0.064
	(0.004)	(0.005)	(0.018)	(0.006)
Effect of process innovation	0.088	0.082	0.102	0.089
	(0.010)	(0.008)	(0.012)	(0.008)
Residuals	-0.003	-0.009	0.072	0.009
	(0.003)	(0.006)	(0.011)	(0.003)

Notes: The contribution of the computer industry to the aggregate TFP growth is calculated from a weighted sum of the TFP growth rates of the four computer industries with the Domar weights. Domar weight is defined as the ratio of gross output in the computer industry to value-added in the non-farm business sector. Numbers in parentheses are the standard errors that are evaluated at their mean values and are also obtained by using the bootstrap method with replacement 1000 times.

<sup>a</sup> : percent per year.

<sup>b</sup> : includes the period of 1974-1990.

<sup>c</sup> : includes the period of 1974-1999.